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Prediction in a Generalized Spatial Panel Data Model with Serial Correlation

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Prediction in a Generalized Spatial Panel Data Model with Serial Correlation

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Abstract

This paper considers the generalized spatial panel data model with serial correlation proposed by Lee and Yu (2012) which encompasses a lot of the spatial panel data models considered in the literature, and derives the best linear unbiased predictor (BLUP) for that model. This in turn provides valuable BLUP for several spatial panel models as special cases.

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Key Words: Prediction; Panel Data; Fixed Effects; Random Effects; Serial Correlation; Spatial Error Correlation

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1 Introduction

Panel data has been used in forecasting gasoline demand across OECD countries, see Baltagi and Griffin (1997); Residential electricity and natural-gas demand using a panel of American states, see Maddala, Trost, Li and Joutz (1997); World carbon dioxide emissions, see Schmalensee, Stoker and Judson (1998); Growth rates of OECD countries, see Hoogstrate, Palm and Pfann (2000); Cigarette sales using a panel of American states, see Baltagi and Li (2004); The impact of uncertainty on U.K. investment authorizations using a panel of U.K. industries, see Driver, Imai, Temple and Urga (2004); Sale of state lottery tickets using panel data on postal (ZIP) codes, see Frees and Miller (2004); Exchange rate determination using industrialized countries quarterly panel data, see Rapach and Wohar (2004); Migration to Germany from 18 source countries over the period 1967-2001, see Brucker and Siliverstovs (2006); Short-term forecasts of employment in a panel of 326 West German regional labor markets observed over the period 1987-2002, see Longhi and Nijkamp (2007); Annual growth rates of real gross regional product for a panel of Chinese regions, see Girardin and Kholodilin (2011), to mention a few. See Baltagi (2013) for a summary of selected empirical panel data forecasting applications.

Wansbeek and Kapteyn (1978), Lee and Griffiths (1979), and Taub (1979) were among the first contributions in econometrics to the problem of prediction in an error component panel data model. Baltagi and Li (1992) extended this prediction to the case of an error component panel model with serial correlation in the remainder disturbance term. While Baltagi and Li (2004, 2006) extended it to the case of spatial autocorrelation in the remainder disturbance term, and Baltagi, Bresson and Pirotte (2012) carried out an extensive Monte Carlo study comparing forecasts in a spatial panel data model. See Baltagi (2013) for a recent survey in the Handbook of Forecasting. This paper considers the generalized spatial panel data model with serial correlation proposed by Lee and Yu (2012) which encompasses a lot of the spatial panel data models considered in the literature, and derives the best linear unbiased predictor (BLUP) for that model. This in turn provides valuable BLUP for several spatial panel models as special cases.

Section 2 gives a brief description of the Lee and Yu (2012) generalized spatial panel data regression model with serial correlation, while Section 3 derives the best linear unbiased predictor (BLUP) for that model. The Lee and Yu (2012) model encompasses a lot of the spatial and panel regression models used in empirical economics. The BLUP for these special cases are shown to follow easily from our BLUP derivation for the generalized model.

2 The Model

Lee and Yu (2012) considered the following generalized spatial panel data regression model with serial correlation, spatial autocorrelation and random effects:

$$y_{it} = x'_{it}\beta + u_{it}, \quad i = 1, \dots, N; t = 1, \dots, T, \quad (1)$$

where y_{it} is the observation on the i th region for the t th time period, x_{it} denotes the $k \times 1$ vector of observations on the nonstochastic regressors and u_{it} is the regression disturbance. In vector form, the disturbance vector of Equation (1) is assumed to have random region effects, spatially autocorrelated residual disturbances and a first-order autoregressive remainder disturbance term:

$$u_t = u_1 + u_{2t}, \quad (2)$$

with

$$u_1 = \lambda_1 W_1 u_1 + (I_N + \delta_1 M_1) \mu \quad (3)$$

$$u_{2t} = \lambda_2 W_2 u_{2t} + (I_N + \delta_2 M_2) \nu_t \quad (4)$$

and

$$\nu_t = \rho\nu_{t-1} + e_t, \quad (5)$$

where $u'_t = (u_{t1}, \dots, u_{tN})$ and ε_t, ν_t and e_t are similarly defined. $\eta' = (\eta_1, \dots, \eta_N)$ denote the vector of random region effects and $\mu' = (\mu_1, \dots, \mu_N)$ are assumed to be $IIN(0, \sigma_\mu^2)$. λ_1 and λ_2 are the scalar spatial autoregressive coefficients with $|\lambda_1| < 1, |\lambda_2| < 1$, δ_1 and δ_2 are the scalar spatial moving average coefficients with $|\delta_1| < 1, |\delta_2| < 1$, while ρ is the time-wise serial correlation coefficient satisfying $|\rho| < 1$. Following Baltagi, Bresson and Pirotte (2012), we define $B_1 = I_N - \lambda_1 W_1$, $B_2 = I_N - \lambda_2 W_2$, $D_1 = I_N + \delta_1 M_1$ and $D_2 = I_N + \delta_2 M_2$. Equations (3) and (4) can be rewritten as:

$$u_1 = A_1^{-1} \mu, \quad (6)$$

$$u_{2t} = A_2^{-1} \nu_t, \quad (7)$$

where $A_1 = D_1^{-1} B_1$ and $A_2 = D_2^{-1} B_2$.

Following Lee and Yu (2012), we employ the following assumptions:

Assumption 1 W_1, W_2, M_1 and M_2 are nonstochastic spatial weights matrices with zero diagonal elements.

Assumption 2 The disturbances e_{it} , $i = 1, 2, \dots, n$ and $t = 2, 3, \dots, T$, are i.i.d. across i and t with zero mean, variance σ_e^2 , and $E|e_{it}|^{4+\kappa} < \infty$ for some $\kappa > 0$; also, they are independent with $\nu_t \sim (0, \sigma_e^2 / (1 - \rho^2) I_N)$.

Assumption 3 B_1, B_2, D_1 and D_2 are invertible for all $\lambda_1 \in \Lambda_1, \lambda_2 \in \Lambda_2, \delta_1 \in \Delta_1, \delta_2 \in \Delta_2$ and $\rho \in \mathbb{P}$, where $\Lambda_1, \Lambda_2, \Delta_1, \Delta_2$ are compact intervals and \mathbb{P} is a compact subset in $(-1, 1)$. Furthermore, $\lambda_1, \lambda_2, \delta_1, \delta_2$ and ρ are, respectively, in the interiors of $\Lambda_1, \Lambda_2, \Delta_1, \Delta_2$ and \mathbb{P} .

Assumption 4 W_1, W_2, M_1 and M_2 are uniformly bounded in both row and column sums in absolute value (for short, UB). Also, $B_1^{-1}, B_2^{-1}, D_1^{-1}$ and D_2^{-1} are UB, uniformly in $\lambda_1 \in \Lambda_1, \lambda_2 \in \Lambda_2, \delta_1 \in \Delta_1$ and $\delta_2 \in \Delta_2$.

Assumption 5 N is large, whereas T is finite.

Assumption 6 $\mu \sim (0, \sigma_\mu^2 I_N)$ is independent of $(\sqrt{1 - \rho^2} \nu'_1, e'_2, \dots, e'_T)'$. Both of them are i.i.d. and independent of X .

As pointed out by Lee and Yu (2012), this model nests various spatial panel models in the literature including the following¹:

¹See Lee and Yu (2010, 2015) for nice surveys of spatial panel data models.

1. When $\lambda_1 = 0$ and $\delta_1 = \delta_2 = 0$, the model reduces to the random effects spatial autoregressive RE-SAR model with serial correlation in the remainder disturbances considered by Baltagi, Song, Jung and Koh (2007).
2. When $\lambda_1 = \lambda_2 = 0$ and $\delta_1 = \delta_2 = 0$, the model reduces to the random effects panel data model with AR(1) remainder error term and no spatial correlation considered by Baltagi and Li (1992).
3. When $\lambda_1 = 0$, $\delta_1 = \delta_2 = 0$ and $\rho = 0$, the model reduces to the random effects spatial autoregressive RE-SAR model with no serial correlation considered by Anselin (1988).
4. When $\lambda_1 = \lambda_2 = 0$, $\delta_1 = 0$ and $\rho = 0$, the model reduces to the random effects spatial moving average RE-SMA model with no serial correlation described by Anselin, Le Gallo and Jayet (2008).
5. When $\lambda_1 = \lambda_2$, $\delta_1 = \delta_2 = 0$, $\rho = 0$ and $W_1 = W_2$, the model reduces to the spatial autoregressive random effects SAR-RE model with no serial correlation considered by Kapoor, Kelejian and Prucha (2007).
6. When $\lambda_1 = \lambda_2 = 0$, $\delta_1 = \delta_2$, $\rho = 0$ and $M_1 = M_2$, the model reduces to the spatial moving average random effects SMA-RE model with no serial correlation considered by Fingleton (2008).
7. When $\delta_1 = \delta_2 = 0$ and $\rho = 0$, the model reduces to the generalized random effects spatial autoregressive model proposed by Baltagi, Egger and Pfaffermayr (2013).
8. When $\lambda_1 = \lambda_2 = 0$, $\delta_1 = \delta_2 = 0$ and $\rho = 0$, the model reduces to the familiar random effects (RE) panel data model with no spatial effects and no serial correlation.

The model in Equation (1) can be rewritten in matrix notation as

$$y = X\beta + u, \quad (8)$$

where y is of dimension $NT \times 1$, X is $NT \times k$, β is $k \times 1$ and u is $NT \times 1$. X is assumed to be of full column rank and its elements are assumed to be bounded in absolute value. The disturbance term can be written in vector form as

$$u = (\iota_T \otimes A_1^{-1}) \mu + (I_T \otimes A_2^{-1}) \nu, \quad (9)$$

where $v' = (v'_1, \dots, v'_T)$ and u is similarly defined. ι_T is a vector of ones of dimension T . I_T is an identity matrix of dimension T and \otimes denotes the Kronecker product.

Under the random effects model, Lee and Yu (2012) showed that the variance–covariance matrix of u can be written as

$$\Omega = E(uu') = \sigma_\mu^2 J_T \otimes (A_1' A_1)^{-1} + \sigma_e^2 V \otimes (A_2' A_2)^{-1}, \quad (10)$$

where J_T is a matrix of ones of dimension T , and $E(vv') = \sigma_e^2 V \otimes I_N$, where V is the familiar AR(1) variance-covariance matrix of dimension T , i.e.,

$$\mathbf{V} = \frac{1}{1 - \rho^2} \begin{bmatrix} 1 & \rho & \rho^2 & \cdots & \rho^{T-1} \\ \rho & 1 & \rho & \cdots & \rho^{T-2} \\ \rho^2 & \rho & 1 & \cdots & \rho^{T-3} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \rho^{T-1} & \rho^{T-2} & \rho^{T-3} & \cdots & 1 \end{bmatrix}. \quad (11)$$

One can easily verify that $\mathbf{V}^{-1} = \mathbf{C}'\mathbf{C}$, where

$$\mathbf{C} = \begin{bmatrix} \sqrt{1 - \rho^2} & 0 & 0 & \cdots & 0 & 0 \\ -\rho & 1 & 0 & \cdots & 0 & 0 \\ 0 & -\rho & 1 & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & -\rho & 1 & 0 \\ 0 & 0 & 0 & 0 & -\rho & 1 \end{bmatrix} \quad (12)$$

is the Prais-Winsten transformation matrix as in Baltagi and Li (1992). From Equation (9), the transformed spatial panel data regression disturbances are given by

$$u^* = (C \otimes I_N) u = (C \iota_T \otimes A_1^{-1}) \mu + (C \otimes A_2^{-1}) v = (1 - \rho) (\iota_T^\alpha \otimes A_1^{-1}) \mu + (C \otimes A_2^{-1}) v, \quad (13)$$

where $C \iota_T = (1 - \rho) \iota_T^\alpha$, where $\iota_T^\alpha = (\alpha, \iota_{T-1}')$ and $\alpha = \sqrt{(1 + \rho)/(1 - \rho)}$. Therefore, the variance–covariance matrix of the Prais–Winsten-transformed spatial panel data model is given by

$$\Omega^* = E(u^* u^{*'}) = (1 - \rho)^2 \sigma_\mu^2 \iota_T^\alpha \iota_T^{\alpha'} \otimes (A_1' A_1)^{-1} + \sigma_e^2 I_T \otimes (A_2' A_2)^{-1} \quad (14)$$

since $(C \otimes A_2^{-1}) E(vv') (C \otimes A_2^{-1})' = \sigma_e^2 I_T \otimes (A_2' A_2)^{-1}$. Replace $\iota_T^\alpha \iota_T^{\alpha'}$ by its idempotent counterpart $d^2 \bar{J}_T^\alpha$, where $\bar{J}_T^\alpha = \iota_T^\alpha \iota_T^{\alpha'} / d^2$ and $d^2 = \iota_T^{\alpha'} \iota_T^\alpha = \alpha^2 + T - 1$. Replace I_T by $E_T^\alpha + \bar{J}_T^\alpha$, where $E_T^\alpha = I_T - \bar{J}_T^\alpha$, and collect like terms, see Baltagi and Li (1992), we get

$$\Omega^* = E(u^* u^{*'}) = \bar{J}_T^\alpha \otimes \left[d^2 (1 - \rho)^2 \sigma_\mu^2 (A_1' A_1)^{-1} + \sigma_e^2 (A_2' A_2)^{-1} \right] + \sigma_e^2 E_T^\alpha \otimes (A_2' A_2)^{-1}. \quad (15)$$

Hence we have

$$\Omega^{*-1} = \bar{J}_T^\alpha \otimes Z + \sigma_e^{-2} E_T^\alpha \otimes A_2' A_2. \quad (16)$$

where $Z = \left[d^2 (1 - \rho)^2 \sigma_\mu^2 (A_1' A_1)^{-1} + \sigma_e^2 (A_2' A_2)^{-1} \right]^{-1}$. Note that Ω in Equation (10) is related to Ω^* in Equation (14) by $\Omega^* = (C \otimes I_N)' \Omega (C \otimes I_N)$. Therefore,

$$\Omega^{-1} = (C \otimes I_N)' \Omega^{*-1} (C \otimes I_N) = (C \otimes I_N)' (\bar{J}_T^\alpha \otimes Z + \sigma_e^{-2} E_T^\alpha \otimes A_2' A_2) (C \otimes I_N). \quad (17)$$

One can easily verify that Equation (17) is equivalent to the inverse of the variance-covariance matrix given by Lee and Yu (2012)

$$\Omega^{-1} = \frac{1}{d^2 (1 - \rho)^2} (V^{-1} \iota_T \iota_T' V^{-1} \otimes Z) + \sigma_e^{-2} \left[\left(V^{-1} - \frac{1}{d^2 (1 - \rho)^2} V^{-1} \iota_T \iota_T' V^{-1} \right) \otimes A_2' A_2 \right] \quad (18)$$

using

$$\frac{1}{d^2 (1 - \rho)^2} V^{-1} \iota_T \iota_T' V^{-1} = \frac{1}{d^2 (1 - \rho)^2} C' C \iota_T \iota_T' C' C = \frac{1}{d^2} C' \iota_T^\alpha \iota_T^{\alpha'} C = C' \bar{J}_T^\alpha C$$

and

$$V^{-1} - \frac{1}{d^2 (1 - \rho)^2} V^{-1} \iota_T \iota_T' V^{-1} = C' C - C' \bar{J}_T^\alpha C = C' E_T^\alpha C.$$

Note that $|\Omega^*| = \left| d^2 (1 - \rho)^2 \sigma_\mu^2 (A_1' A_1)^{-1} + \sigma_e^2 (A_2' A_2)^{-1} \right| \left| \sigma_e^2 (A_2' A_2)^{-1} \right|^{T-1}$, $|C| = \sqrt{1 - \rho^2}$ and $|C \otimes I_N| = |C|^N$, see Magnus (1982). Under the assumption of normality, the log-likelihood function for this model can be written as

$$\begin{aligned} L = & \text{Const.} + \frac{1}{2} N \ln (1 - \rho^2) - \frac{1}{2} \ln \left| d^2 (1 - \rho)^2 \sigma_\mu^2 (A_1' A_1)^{-1} + \sigma_e^2 (A_2' A_2)^{-1} \right| \\ & - \frac{1}{2} N (T - 1) \ln \sigma_e^2 + (T - 1) \ln |A_2| - \frac{1}{2} u^{*'} \Omega^{*-1} u^*, \end{aligned} \quad (19)$$

where u^* is given by Equation (13) and Ω^{*-1} is given by Equation (16).

Assumption 7 Elements of the $N \times k$ matrix of regressors X are nonstochastic and bounded, uniformly in N and T . Also, under the asymptotic setting in Assumption 5, the limit of $\frac{1}{NT} \sum_{t=1}^T X' \Omega^{-1} X$ exists and is nonsingular.

Assumption 8 $\lim_{N,T \rightarrow \infty} \left[\frac{1}{NT} \ln |\Omega| + 1 - \left(\frac{1}{NT} \ln |\Omega(\phi)| + p_{NT}(\phi) \right) \right] \neq 0$ for $\phi \neq \phi_0$, where $p_{NT}(\phi) = \frac{1}{NT} \text{tr} |\Omega^{-1}(\phi) \Omega|$, $\phi = (\lambda_1, \lambda_2, \delta_1, \delta_2, \rho, \sigma_\mu^2, \sigma_e^2)$ and ϕ_0 denotes the true value of ϕ .

Under Assumptions 1-8, Lee and Yu (2012) establishes consistency and asymptotic normality of the quasi-maximum likelihood estimator. They provided Matlab programs for these estimation methods. See

also Millo (2014) for R programs performing maximum likelihood estimation of panel data models with random effects, a spatially lagged dependent variable and spatially and serially correlated errors. In this paper we are interested in prediction. This is taken up in the next section.

3 BLUP

Goldberger (1962) showed that, for a known Ω , the best linear unbiased predictor (BLUP) for the i th individual s periods ahead ($y_{i,T+s}$) is given by

$$\hat{y}_{i,T+s} = x'_{i,T+s} \hat{\beta}_{GLS} + w' \Omega^{-1} \hat{u}_{GLS}, \quad (20)$$

where $w = E(uu_{i,T+s})$ is the covariance between the future disturbance $u_{i,T+s}$ and the sample disturbances u . $\hat{\beta}_{GLS}$ is the GLS estimator of β from Equation (8) based on the true Ω . Also, $\hat{u}_{GLS} = y - x' \hat{\beta}_{GLS}$ denotes the corresponding GLS residual vector. From Equation (9), $u_{i,T+s}$ can be rewritten as $u_{i,T+s} = \eta_i + \varepsilon_{i,T+s} = l'_i A_1^{-1} \mu + l'_i A_2^{-1} v_{T+s}$, where l'_i as the i th row of I_N and v_{T+s} is the $N \times 1$ vector of disturbances for the $(T+s)$ th time period. Focusing on the last term of Equation (20), which we will call the Goldberger BLUP term, we get

$$w' \Omega^{-1} \hat{u}_{GLS} = E(u_{i,T+s} u') \Omega^{-1} \hat{u}_{GLS} = E(l'_i A_1^{-1} \mu u') \Omega^{-1} \hat{u}_{GLS} + E(l'_i A_2^{-1} v_{T+s} u') \Omega^{-1} \hat{u}_{GLS}. \quad (21)$$

Consider the first term in Equation (21). Define $Z_1 = (A'_1 A_1)^{-1} Z$. Using Equation (13), The first term in Equation (21) can be expressed as:

$$\begin{aligned} & E(l'_i A_1^{-1} \mu u') \Omega^{-1} \hat{u}_{GLS} = E(l'_i A_1^{-1} \mu u') (C \otimes I_N)' \Omega^{*-1} (C \otimes I_N) \hat{u}_{GLS} \\ &= E \left\{ l'_i A_1^{-1} \mu [(1-\rho)(\iota_T^\alpha \otimes A_1^{-1}) \mu + (C \otimes A_2^{-1}) v]' \right\} \Omega^{*-1} \hat{u}_{GLS}^* \\ &= l'_i A_1^{-1} E(\mu \mu') (1-\rho)(\iota_T^\alpha \otimes A_1^{-1})' \Omega^{*-1} \hat{u}_{GLS}^* \\ &= (1-\rho) \sigma_\mu^2 \left[\iota_T^{\alpha'} \otimes l'_i (A'_1 A_1)^{-1} \right] [\bar{J}_T^\alpha \otimes Z + \sigma_e^{-2} E_T^\alpha \otimes A'_2 A_2] \hat{u}_{GLS}^* \\ &= (1-\rho) \sigma_\mu^2 \left[\iota_T^{\alpha'} \otimes l'_i (A'_1 A_1)^{-1} Z \right] \hat{u}_{GLS}^* \\ &= (1-\rho) \sigma_\mu^2 (\iota_T^{\alpha'} \otimes l'_i Z_1) \hat{u}_{GLS}^* \\ &= (1-\rho) \sigma_\mu^2 \sum_{k=1}^N \left[z_{1ik} \left(\alpha \hat{u}_{k1}^* + \sum_{t=2}^T \hat{u}_{kt}^* \right) \right], \end{aligned} \quad (22)$$

where z_{1ik} is the (i, k) th elements of Z_1 and \hat{u}_{it}^* is the it th elements of $\hat{u}_{GLS}^* = (C \otimes I_N) \hat{u}_{GLS}$. This uses the following results: $\iota_T^{\alpha'} \bar{J}_T^\alpha = \iota_T^{\alpha'}$, $\iota_T^{\alpha'} E_T^\alpha = \mathbf{0}$ and μ and v_t are independent.

Consider the second term in Equation (21). Notice that

$$\begin{aligned}
E(l'_i A_2^{-1} v_{T+s} u') &= E\left\{l'_i A_2^{-1} v_{T+s} [(\iota_T \otimes A_1^{-1}) \mu + (I_T \otimes A_2^{-1}) \nu]'\right\} \\
&= l'_i A_2^{-1} E(v_{T+s} \nu') (I_T \otimes A_2^{-1})' \\
&= l'_i A_2^{-1} \left[\frac{\sigma_e^2}{1-\rho^2} \rho^s (\rho^{T-1}, \rho^{T-2}, \dots, \rho, 1) \otimes I_N \right] (I_T \otimes A_2^{-1})' \\
&= \frac{\sigma_e^2}{1-\rho^2} \rho^s \left[(\rho^{T-1}, \rho^{T-2}, \dots, \rho, 1) \otimes l'_i (A'_2 A_2)^{-1} \right]
\end{aligned} \tag{23}$$

since μ and v_t are independent, and Ω^{-1} in Equation (18) can be rewritten as

$$\begin{aligned}
\Omega^{-1} &= \frac{1}{d^2 (1-\rho)^2} (V^{-1} \iota_T \iota_T' V^{-1} \otimes Z) + \sigma_e^{-2} \left[\left(V^{-1} - \frac{1}{d^2 (1-\rho)^2} V^{-1} \iota_T \iota_T' V^{-1} \right) \otimes A'_2 A_2 \right] \\
&= \sigma_e^{-2} V^{-1} \otimes A'_2 A_2 + \frac{1}{d^2 (1-\rho)^2} (V^{-1} \iota_T \iota_T' V^{-1}) \otimes (Z - \sigma_e^{-2} A'_2 A_2) \\
&= (\sigma_e^{-2} V^{-1} \otimes A'_2 A_2) \left[I_{TN} + \frac{1}{d^2 (1-\rho)} (\iota_T \iota_T' C) \otimes (\sigma_e^2 Z_2 - I_N) \right],
\end{aligned} \tag{24}$$

where $Z_2 = (A'_2 A_2)^{-1} Z$. Also, $\iota_T' V^{-1} = \iota_T' C' C = (1-\rho) \iota_T' C$. Hence the second term in Equation (21) can be written as:

$$\begin{aligned}
&E(l'_i A_2^{-1} v_{T+s} u') \Omega^{-1} \hat{u}_{GLS} \\
&= \frac{\sigma_e^2}{1-\rho^2} \rho^s \left[(\rho^{T-1}, \rho^{T-2}, \dots, \rho, 1) \otimes l'_i (A'_2 A_2)^{-1} \right] (\sigma_e^{-2} V^{-1} \otimes A'_2 A_2) \\
&\quad \left[I_{TN} + \frac{1}{d^2 (1-\rho)} (\iota_T \iota_T' C) \otimes (\sigma_e^2 Z_2 - I_N) \right] \hat{u}_{GLS} \\
&= [\rho^s (0, 0, \dots, 0, 1) \otimes l'_i] \left\{ \hat{u}_{GLS} + \frac{1}{d^2 (1-\rho)} [(\iota_T \iota_T' C) \otimes (\sigma_e^2 Z_2 - I_N)] \hat{u}_{GLS}^* \right\} \\
&= \rho^s [(0, 0, \dots, 0, 1) \otimes l'_i] \hat{u}_{GLS} + \frac{\rho^s}{d^2 (1-\rho)} [l'_i C \otimes (\sigma_e^2 l'_i Z_2 - l'_i)] \hat{u}_{GLS}^* \\
&= \rho^s \hat{u}_{i,T} + \frac{\sigma_e^2 \rho^s}{d^2 (1-\rho)} \sum_{k=1}^N \left[z_{2ik} \left(\alpha \hat{u}_{k1}^* + \sum_{t=2}^T \hat{u}_{kt}^* \right) \right] - \frac{\rho^s}{d^2 (1-\rho)} \left(\alpha \hat{u}_{i1}^* + \sum_{t=2}^T \hat{u}_{it}^* \right),
\end{aligned} \tag{25}$$

where z_{2ik} is the (i, k) th element of Z_2 and $\hat{u}_{GLS}^* = (C \otimes I_N) \hat{u}_{GLS}$. This uses the following results: $\frac{1}{1-\rho^2} (\rho^{T-1}, \rho^{T-2}, \dots, \rho, 1)$ is the last row of V and $(0, 0, \dots, 0, 1) \iota_T = 1$. Combining Equations (22) and

(25), one gets the following Goldberger BLUP term:

$$\begin{aligned}
w'\Omega^{-1}\hat{u}_{GLS} &= (1-\rho)\sigma_\mu^2\sum_{k=1}^N\left[z_{1ik}\left(\alpha\hat{u}_{k1}^*+\sum_{t=2}^T\hat{u}_{kt}^*\right)\right] \\
&\quad +\rho^s\hat{u}_{i,T}+\frac{\sigma_e^2\rho^s}{d^2(1-\rho)}\sum_{k=1}^N\left[z_{2ik}\left(\alpha\hat{u}_{k1}^*+\sum_{t=2}^T\hat{u}_{kt}^*\right)\right]-\frac{\rho^s}{d^2(1-\rho)}\left(\alpha\hat{u}_{i1}^*+\sum_{t=2}^T\hat{u}_{it}^*\right) \\
&= \rho^s\hat{u}_{i,T}+\sum_{k=1}^N\left[\left((1-\rho)\sigma_\mu^2z_{1ik}+\frac{\rho^s\sigma_e^2}{d^2(1-\rho)}z_{2ik}\right)\left(\alpha\hat{u}_{k1}^*+\sum_{t=2}^T\hat{u}_{kt}^*\right)\right] \\
&\quad -\frac{\rho^s}{d^2(1-\rho)}\left(\alpha\hat{u}_{i1}^*+\sum_{t=2}^T\hat{u}_{it}^*\right). \tag{26}
\end{aligned}$$

Special case 1: When $\lambda_1 = 0$ and $\delta_1 = \delta_2 = 0$, the model reduces to the random effects spatial autoregressive RE-SAR model *with serial correlation* considered by Baltagi, Song, Jung and Koh (2007). In this case, we have $A_1 = I_N$, $A_2 = B_2$, $Z = \left[d^2(1-\rho)^2\sigma_\mu^2I_N + \sigma_e^2(B_2'B_2)^{-1}\right]^{-1}$, $Z_1 = Z$ and $Z_2 = (B_2'B_2)^{-1}Z$. The Goldberger BLUP term given in Equation (26) reduces to

$$\begin{aligned}
w'\Omega^{-1}\hat{u}_{GLS} &= \rho^s\hat{u}_{i,T}+\sum_{k=1}^N\left[\left((1-\rho)\sigma_\mu^2z_{ik}+\frac{\rho^s\sigma_e^2}{d^2(1-\rho)}g_{ik}\right)\left(\alpha\hat{u}_{k1}^*+\sum_{t=2}^T\hat{u}_{kt}^*\right)\right] \\
&\quad -\frac{\rho^s}{d^2(1-\rho)}\left(\alpha\hat{u}_{i1}^*+\sum_{t=2}^T\hat{u}_{it}^*\right), \tag{27}
\end{aligned}$$

where z_{ik} and g_{ik} are the (i, k) th elements of Z and $(B_2'B_2)^{-1}Z$, respectively. Equivalently, g_{ik} can be defined as the k th element of $g'_i = b'_iB_2^{-1}Z$ or $g_i = Z'B_2^{-1}b_i$, where b'_i as the i th row of B_2^{-1} . This is Goldberger's BLUP *extra term* derived by Song and Jung (2002) for the random effects error component model with SAR correlation and serial correlation in the remainder disturbances.

Special case 2: When $\lambda_1 = \lambda_2 = 0$ and $\delta_1 = \delta_2 = 0$, the model reduces to the random effects panel data model with AR(1) remainder error term and *no spatial correlation* considered by Baltagi and Li (1992). This model is special case 1, but with no spatial correlation. In this case, we have $A_1 = A_2 = I_N$, $Z = Z_1 = Z_2 = \left[d^2(1-\rho)^2\sigma_\mu^2I_N + \sigma_e^2I_N\right]^{-1} = \sigma_\alpha^{-2}I_N$, where $\sigma_\alpha^2 = d^2(1-\rho)^2\sigma_\mu^2 + \sigma_e^2$. Substituting these

terms into Equation (27), the Goldberger BLUP term given in Equation (26) reduces to

$$\begin{aligned}
w'\Omega^{-1}\hat{u}_{GLS} &= \rho^s \hat{u}_{i,T} + \sum_{k=1}^N \left[\left(\frac{(1-\rho)\sigma_\mu^2}{\sigma_\alpha^2} l_{ik} + \frac{\rho^s \sigma_e^2}{d^2(1-\rho)\sigma_\alpha^2} l_{ik} \right) \left(\alpha \hat{u}_{k1}^* + \sum_{t=2}^T \hat{u}_{kt}^* \right) \right. \\
&\quad \left. - \frac{\rho^s}{d^2(1-\rho)} \left(\alpha \hat{u}_{i1}^* + \sum_{t=2}^T \hat{u}_{it}^* \right) \right] \\
&= \rho^s \hat{u}_{i,T} + \left(\frac{(1-\rho)\sigma_\mu^2}{\sigma_\alpha^2} + \frac{\rho^s \sigma_e^2}{d^2(1-\rho)\sigma_\alpha^2} \right) \left(\alpha \hat{u}_{i1}^* + \sum_{t=2}^T \hat{u}_{it}^* \right) \\
&\quad - \frac{\rho^s}{d^2(1-\rho)} \left(\alpha \hat{u}_{i1}^* + \sum_{t=2}^T \hat{u}_{it}^* \right) \\
&= \rho^s \hat{u}_{i,T} + \frac{(1-\rho^s)(1-\rho)\sigma_\mu^2}{\sigma_\alpha^2} \left(\alpha \hat{u}_{i1}^* + \sum_{t=2}^T \hat{u}_{it}^* \right), \tag{28}
\end{aligned}$$

where l_{ik} is the (i, k) th elements of I_N . When $s = 1$, it further reduces to Goldberger's BLUP *extra term* derived by Baltagi and Li (1992) for the random effects panel data model with AR(1) remainder error term and no spatial correlation.

Special case 3: When $\lambda_1 = 0$, $\delta_1 = \delta_2 = 0$ and $\rho = 0$, the model reduces to the random effects spatial autoregressive RE-SAR model with no serial correlation considered by Anselin (1988). This is special case 1, but with no serial correlation. Note that $\rho = 0$ implies that $\alpha = 1$, $d^2 = T$, $\hat{u}_{GLS}^* = \hat{u}_{GLS}$ and $Z = C_1 \equiv [T\sigma_\mu^2 I_N + \sigma_e^2 (B_2' B_2)^{-1}]^{-1}$. Substituting these into Equation (27), the Goldberger BLUP term given in Equation (26) reduces to

$$w'\Omega^{-1}\hat{u}_{GLS} = \sum_{k=1}^N \left(\sigma_\mu^2 c_{1ik} \sum_{t=1}^T \hat{u}_{kt} \right) = T\sigma_\mu^2 \sum_{k=1}^N (c_{1ik} \bar{u}_k), \tag{29}$$

where c_{1ik} is the (i, k) th elements of C_1 and $\bar{u}_k = \frac{1}{T} \sum_{t=1}^T \hat{u}_{kt}$. This is Goldberger's BLUP *extra term* derived by Baltagi and Li (2004, 2006) for the random effects error component model with SAR correlation in the remainder disturbances.

Special case 4: When $\lambda_1 = \lambda_2 = 0$, $\delta_1 = 0$ and $\rho = 0$, the model reduces to the random effects spatial moving average RE-SMA model with *no serial correlation* described by Anselin, Le Gallo and Jayet (2008). In this case, we have $A_1 = I_N$, $A_2 = D_2^{-1}$. Note that $\rho = 0$ implies that $\alpha = 1$, $d^2 = T$, $\hat{u}_{GLS}^* = \hat{u}_{GLS}$ and $Z = C_2 \equiv [T\sigma_\mu^2 I_N + \sigma_e^2 D_2' D_2]^{-1}$, $Z_1 = Z$ and $Z_2 = D_2' D_2 Z$. Substituting these into Equation (27), the Goldberger BLUP term given in Equation (26) reduces to

$$w'\Omega^{-1}\hat{u}_{GLS} = \sum_{k=1}^N \left(\sigma_\mu^2 c_{2ik} \sum_{t=1}^T \hat{u}_{kt} \right) = T\sigma_\mu^2 \sum_{k=1}^N (c_{2ik} \bar{u}_k), \tag{30}$$

where c_{2ik} is the (i, k) th elements of C_2 and $\bar{u}_k = \frac{1}{T} \sum_{t=1}^T \hat{u}_{kt}$. This is Goldberger's BLUP *extra term* derived by Baltagi and Li (2004, 2006) for the RE-SMA model with no serial correlation in the remainder disturbances.

Special case 5: When $\lambda_1 = \lambda_2$, $\delta_1 = \delta_2 = 0$, $W_1 = W_2$ and $\rho = 0$, the model reduces to the spatial autoregressive random effects SAR-RE model with *no serial correlation* considered by Kapoor, Kelejian and Prucha (2007). In this case, we have $A_1 = A_2 = B_1$. Note that $\rho = 0$ implies that $\alpha = 1$, $d^2 = T$, $\hat{u}_{GLS}^* = \hat{u}_{GLS}$ and $Z = \left[T\sigma_\mu^2 (B_1' B_1)^{-1} + \sigma_e^2 (B_1' B_1)^{-1} \right]^{-1} = \sigma_1^{-2} B_1' B_1$, where $\sigma_1^2 = T\sigma_\mu^2 + \sigma_e^2$ and $Z_1 = Z_2 = \sigma_1^{-2} I_N$. Substituting these results into Equation (27), the Goldberger BLUP term given in Equation (26) reduces to

$$w' \Omega^{-1} \hat{u}_{GLS} = \sum_{k=1}^N \left(\frac{\sigma_\mu^2}{\sigma_1^2} l_{ik} \sum_{t=1}^T \hat{u}_{kt} \right) = \frac{T\sigma_\mu^2}{\sigma_1^2} \bar{u}_i, \quad (31)$$

where l_{ik} is the (i, k) th elements of I_N and $\bar{u}_i = \frac{1}{T} \sum_{t=1}^T \hat{u}_{it}$. This is equivalent to $\frac{\sigma_\mu^2}{\sigma_1^2} (l'_T \otimes l'_i) \hat{u}_{GLS}$, where l'_i as the i th row of I_N . This is Goldberger's BLUP *extra term* derived by Baltagi, Bresson and Pirotte (2012) for the SAR-RE model with no serial correlation in the remainder disturbances.

Special case 6: When $\lambda_1 = \lambda_2 = 0$, $\delta_1 = \delta_2$, $M_1 = M_2$ and $\rho = 0$, the model reduces to the spatial moving average random effects SMA-RE model with *no serial correlation* considered by Fingleton (2008). In this case, we have $A_1 = A_2 = D_2^{-1} \equiv (I_N - \delta_2 M_2)^{-1}$. Note that $\rho = 0$ implies that $\alpha = 1$, $d^2 = T$, $\hat{u}_{GLS}^* = \hat{u}_{GLS}$ and $Z = [T\sigma_\mu^2 D_2' D_2 + \sigma_e^2 D_2' D_2]^{-1} = \sigma_1^{-2} (D_2' D_2)^{-1}$, where $\sigma_1^2 = T\sigma_\mu^2 + \sigma_e^2$ and $Z_1 = Z_2 = \sigma_1^{-2} I_N$. Substituting these results into Equation (27), the Goldberger BLUP term given in Equation (26) reduces to

$$w' \Omega^{-1} \hat{u}_{GLS} = \sum_{k=1}^N \left(\frac{\sigma_\mu^2}{\sigma_1^2} l_{ik} \sum_{t=1}^T \hat{u}_{kt} \right) = \frac{T\sigma_\mu^2}{\sigma_1^2} \bar{u}_i, \quad (32)$$

where l_{ik} is the (i, k) th elements of I_N and $\bar{u}_i = \frac{1}{T} \sum_{t=1}^T \hat{u}_{it}$. This is again equivalent to $\frac{\sigma_\mu^2}{\sigma_1^2} (l'_T \otimes l'_i) \hat{u}_{GLS}$, where l'_i as the i th row of I_N . This is Goldberger's BLUP *extra term* derived by Baltagi, Bresson and Pirotte (2012) for the SMA-RE model with no serial correlation in the remainder disturbances and it is the same as the one for SAR-RE model with no serial correlation in the remainder disturbances considered in special case 5. Note, however, that the feasible predictor will be based on different estimates of the residuals and variance components once the model is estimated by maximum likelihood or Generalized Moments.

Special case 7: When $\delta_1 = \delta_2 = 0$ and $\rho = 0$, the model reduces to the generalized random effects spatial autoregressive model with *no serial correlation*, proposed by Baltagi, Egger and Pfaffermayr (2013). In this case, we have $A_1 = B_1$ and $A_2 = B_2$. Note that $\rho = 0$ implies that $\alpha = 1$, $d^2 = T$, $\hat{u}_{GLS}^* = \hat{u}_{GLS}$ and $Z = \left[T\sigma_\mu^2 (B_1' B_1)^{-1} + \sigma_e^2 (B_2' B_2)^{-1} \right]^{-1}$ and $Z_1 = (B_1' B_1)^{-1} Z = \left[T\sigma_\mu^2 I_N + \sigma_e^2 (B_2' B_2)^{-1} (B_1' B_1) \right]^{-1}$

$\equiv C_3$. Substituting these results into Equation (27), the Goldberger BLUP term given in Equation (26) reduces to

$$w'\Omega^{-1}\hat{u}_{GLS} = \sum_{k=1}^N \left(\sigma_\mu^2 c_{3ik} \sum_{t=1}^T \hat{u}_{kt} \right) = T\sigma_\mu^2 \sum_{k=1}^N (c_{3ik} \bar{u}_k), \quad (33)$$

where c_{3ik} is the (i, k) th elements of C_3 and $\bar{u}_k = \frac{1}{T} \sum_{t=1}^T \hat{u}_{kt}$.

Special case 8: When $\lambda_1 = \lambda_2 = 0$, $\delta_1 = \delta_2 = 0$ and $\rho = 0$, the model reduces to the familiar random effects model *without spatial or serial autocorrelation*. In this case, we have $A_1 = A_2 = I_N$. Note that $\rho = 0$ implies that $\alpha = 1$, $d^2 = T$, $\hat{u}_{GLS}^* = \hat{u}_{GLS}$ and $Z = Z_1 = Z_2 = [T\sigma_\mu^2 I_N + \sigma_e^2 I_N]^{-1} = \sigma_1^{-2} I_N$, where $\sigma_1^2 = T\sigma_\mu^2 + \sigma_e^2$. Substituting these results into Equation (27), the Goldberger BLUP term given in Equation (26) reduces to

$$w'\Omega^{-1}\hat{u}_{GLS} = \sum_{k=1}^N \left(\frac{\sigma_\mu^2}{\sigma_1^2} l_{ik} \sum_{t=1}^T \hat{u}_{kt} \right) = \frac{T\sigma_\mu^2}{\sigma_1^2} \bar{u}_i, \quad (34)$$

where l_{ik} is the (i, k) th elements of I_N and $\bar{u}_i = \frac{1}{T} \sum_{t=1}^T \hat{u}_{it}$. This is again equivalent to $\frac{\sigma_\mu^2}{\sigma_1^2} (l'_T \otimes l'_i) \hat{u}_{GLS}$, where l'_i as the i th row of I_N . This is Goldberger's BLUP *extra term* derived by Wansbeek and Kapteyn (1978), Lee and Griffiths (1979), and Taub (1979) for the random effects error component model and it is the same as the one for SAR or SMA correlation in the remainder disturbances in special cases 5 and 6 but with different estimates of the residuals and variance components once the model is estimated by maximum likelihood or Generalized Moments. In order to make this forecast operational, $\hat{\beta}_{GLS}$ is replaced by its feasible GLS estimate and the variance components are replaced by their feasible estimates.

4 Monte Carlo Simulation

This section performs some Monte Carlo experiments to evaluate the performance of our proposed predictors for the random effects model with both time autocorrelated and spatial correlated disturbances. It is important to note that Baltagi, Bresson and Pirotte (2012) performed extensive Monte Carlo experiments to evaluate the performance of predictors for the random effects model with spatial correlated disturbances. Following Baltagi, Bresson and Pirotte (2012) the data generating process starts with a simple panel data regression with random one-way error components disturbances

$$y_{it} = 5 + 0.5x_{it} + u_{it}, \quad i = 1, \dots, N, \quad t = 1, \dots, T + 1. \quad (35)$$

The variable x_{it} was generated as $x_{it} = \delta_i + \xi_{it}$, where δ_i is a random variable uniformly distributed on the interval $[-7.5, 7.5]$ and ξ_{it} is a random variable uniformly distributed on the interval $[-5, 5]$. We choose the same spatial weight matrix $W_1 = W_2 = M_1 = M_2 = W$. Following Baltagi, Bresson and Pirotte (2012), the

matrix W is created such that its i -th row has non-zero elements in positions $i + 5$ and $i - 5$. Therefore, the i -th element of u is directly related to the five ones immediately before it and the five ones immediately after it. This matrix is defined in a circular world so that the non-zero elements in rows 1 and N are, respectively, in positions $(2, 3, 4, 5, 6, N - 4, N - 3, N - 2, N - 1, N)$ and $(1, 2, 3, 4, 5, N - 5, N - 4, N - 3, N - 2, N - 1)$. This matrix is row normalized so that all of its non-zero elements are equal to $1/10$. As in Kapoor, Kelejian and Prucha (2007), this weighting matrix is referred as “5 ahead and 5 behind”. The remainder disturbances u_{it} were generated as an spatially correlated process with the following Data Generating Processes (DGP):

1. SAR: $\delta_1 = \delta_2 = 0$, λ_1 and λ_2 take values $(0, 0.2, 0.5, 0.8)$. These are reported in Tables 1-4 for $\rho = 0, 0.2, 0.5, 0.8$.
2. SMA: $\lambda_1 = \lambda_2 = 0$, δ_1 and δ_2 take values $(0, 0.2, 0.5, 0.8)$. These are reported in Tables 5-8 for $\rho = 0, 0.2, 0.5, 0.8$.
3. SARMA: $\delta_1, \delta_2, \lambda_1$ and λ_2 take values $(0.2, 0.5)$. These are reported in Tables 9-12 for $\rho = 0, 0.2, 0.5, 0.8$.

The individual specific effect μ_i is a random variable uniformly distributed as $\mu_i \stackrel{iid}{\sim} N(0, 10)$. The remainder disturbances ν_{it} were generated as an AR(1) process with $\nu_{it} = \rho\nu_{i,t-1} + \varepsilon_{it}$, where ε_{it} is a random variable uniformly distributed as $\varepsilon_{it} \stackrel{iid}{\sim} N(0, 10)$ and ρ takes values $(0, 0.2, 0.5, 0.8)$. Baltagi, Bresson and Pirotte (2012) considered several forecasts using panel data with spatial error correlation where the true data generating process was assumed to be a simple error component regression model with spatial remainder disturbances of the autoregressive or moving average type. Here, we extend this to the spatial autoregressive moving average type.

Predictions were made for only one period ahead. In order to depict the typical United States panel, the sample sizes (N, T) in the different experiments were chosen as $(49, 10)$. For each experiment, we perform 1,000 replications. For each replication we estimate the model using the first 10 years and forecast 1 year ahead. Following Baltagi, Bresson and Pirotte (2012), we report the sampling root mean square error (RMSE) of each of the predictors considered above, which is computed as

$$RMSE = \frac{1}{R} \sum_{r=1}^R \sqrt{\frac{1}{N} \sum_{i=1}^N (\hat{y}_{i,T+1} - y_{i,T+1})^2}, \quad (36)$$

where $R = 1,000$ replications. Following Frees and Miller (2004) among others, we also summarize the accuracy of the forecasts using the mean absolute error (MAE)

$$MAE = \frac{1}{RN} \sum_{r=1}^R \sum_{i=1}^N |\hat{y}_{i,T+1} - y_{i,T+1}|. \quad (37)$$

For example, Willmott and Matsuura (2005) show that MAE has advantages over RMSE. RMSE and MAE of each of the predictors is reported in Tables 1-12. The columns of these Tables are labeled with the estimator used. The first column is OLS, the second column is the estimator for special case 1 which is a RE-SAR with AR(1) remainder error, the third column is the estimator for special case 2 which is a RE-SAR with no serial correlation, etc. The last column is for the Generalized estimator for a SARMA with AR(1) remainder error. For each model except OLS, slope and variance components parameters are estimated using MLE. It is worth pointing out that MLE estimators from an incorrectly specified model may affect the properties of the forecast. OLS is consistent but not efficient and ignores the heterogeneity in the panel and the spatial correlation. We include it for applied researchers that ignore spatial correlation and heterogeneity in the panel. Obviously, its predictions do not use the Goldberger correction and perform badly in Monte Carlo as the BLUP theory predicts.

Overall, forecasts one year ahead based on OLS, an estimator that ignores heterogeneity, spatial correlation and time autocorrelation performs the worst in terms of RMSE in all Tables. In Tables with $\rho = 0$, predictors one year ahead based on estimators that do not correct for serial correlation perform well in terms of RMSE. As ρ increases to 0.5 and 0.8, predictors one year ahead based on estimators that correct for serial correlation perform well in terms of RMSE. In Table 12, where the DGP is a SARMA with $\rho = 0.8$, the best RMSE is obtained by cases 1, 2, and the General predictor, all of which take care of serial correlation.

Predictors that account for time autocorrelation improve the forecast performance by a big margin. Predictors that account for spatial correlation improve the forecast, but by a smaller margin. These findings are consistent with those in Baltagi, Bresson and Pirotte (2012). This is true whether the true model is SAR, SMA or SARMA with AR(1) remainder error. In Table 4, where the true model is SAR with $\rho = 0.8$, OLS has a RMSE of 8.242 for $\lambda_1 = \lambda_2 = 0.8$. Correcting for heterogeneity using a random effects estimator, as in case 8, only drops this to 6.686. If we do RE correcting for serial and spatial correlation, this forecast RMSE drops to 4.579 for case 1 (SAR-RE) and 4.573 for the General (SARMA-RE) estimator. Note that ignoring the spatial correlation and correcting only for serial correlation as in case 2, RE with AR(1), drops this forecast RMSE already to 4.604. Correcting for spatial correlation without correcting for time wise serial correlation as in cases 3-7 drops this RMSE only to 6.692 to 6.726 range. The results are similar in Table 8, where the true model is SMA with $\rho = 0.8$, OLS has a RMSE of 6.195 for $\delta_1 = \delta_2 = 0.8$. Correcting for heterogeneity using a random effects estimator, as in case 8, only drops this to 4.871. If we do RE correcting for serial and spatial correlation, this forecast RMSE drops to 3.301 for case 1 (SAR-RE) and 3.300 for the General (SARMA-RE) estimator. Note that ignoring the spatial correlation and correcting only for serial correlation as in case 2, RE with AR(1), drops this forecast RMSE already to 3.302. Correcting

for spatial correlation without correcting for time wise serial correlation as in cases 3-7 drops this RMSE only to the 4.867 to 4.873 range. The same thing happens for Table 12, where the true model is SARMA with $\rho = 0.8$, OLS has a RMSE of 7.041 for $\lambda_1 = \lambda_2 = 0.8$ and $\delta_1 = \delta_2 = 0.5$. Correcting for heterogeneity using a random effects estimator, as in case 8, only drops this to 5.598. If we do RE correcting for serial and spatial correlation, this forecast RMSE drops to 3.814 for case 1 (SAR-RE) and 3.810 for the General (SARMA-RE) estimator. Note that ignoring the spatial correlation and correcting only for serial correlation as in case 2, RE with AR(1), drops this forecast RMSE already to 3.821. Correcting for spatial correlation without correcting for time wise serial correlation as in cases 3-7 drops this RMSE only to the 5.591 to 5.601 range. Results of MAE yield similar findings to those of RMSE in Tables 1-12.

5 Conclusion

This paper derives Goldberger's (1962) best linear unbiased predictor (BLUP) for the generalized spatial panel data model with serial correlation proposed by Lee and Yu (2012). Since the latter model encompasses a lot of the spatial panel data models considered in the literature, this in turn provides valuable BLUP for several spatial panel models as special cases. Extensions of this BLUP should be applied to dynamic spatial panel models, see Baltagi, Fingleton and Pirotte (2014), and to panel data models with a spatial lag, as well as higher order autoregressive and moving average processes, see Baltagi and Liu (2013a, 2013b). Furthermore, applied researchers may be interested in confidence intervals for serially dependent data. See Lahiri and Yang (2013) for an example. One might be interested in obtaining confidence intervals of $\hat{y}_{i,T+s}$. This leaves a potential research topic for the future.

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Table 1: RMSE and MAE of Spatial Panel Data Predictors: SAR and $\rho = 0$

RMSE											
λ_1	λ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	4.396	3.290	3.290	3.284	3.284	3.284	3.284	3.285	3.284	3.291
0	0.2	4.417	3.316	3.317	3.311	3.311	3.311	3.311	3.312	3.311	3.319
0	0.5	4.574	3.524	3.529	3.518	3.518	3.519	3.519	3.521	3.519	3.530
0	0.8	5.520	4.684	4.737	4.674	4.678	4.693	4.692	4.680	4.699	4.694
0.2	0	4.399	3.290	3.290	3.285	3.285	3.284	3.284	3.285	3.285	3.291
0.2	0.2	4.420	3.317	3.317	3.312	3.312	3.311	3.311	3.312	3.311	3.318
0.2	0.5	4.576	3.527	3.529	3.521	3.520	3.519	3.519	3.523	3.520	3.530
0.2	0.8	5.521	4.693	4.738	4.684	4.682	4.693	4.692	4.688	4.699	4.702
0.5	0	4.482	3.291	3.291	3.286	3.286	3.286	3.286	3.284	3.286	3.290
0.5	0.2	4.502	3.320	3.318	3.314	3.314	3.312	3.312	3.312	3.312	3.318
0.5	0.5	4.656	3.537	3.531	3.530	3.528	3.520	3.520	3.523	3.521	3.530
0.5	0.8	5.587	4.726	4.740	4.717	4.695	4.694	4.694	4.702	4.702	4.715
0.8	0	4.957	3.295	3.295	3.290	3.290	3.290	3.290	3.283	3.290	3.289
0.8	0.2	4.975	3.326	3.322	3.321	3.320	3.317	3.316	3.311	3.316	3.317
0.8	0.5	5.113	3.561	3.535	3.554	3.543	3.526	3.526	3.523	3.526	3.530
0.8	0.8	5.974	4.886	4.746	4.879	4.726	4.700	4.701	4.708	4.708	4.718
MAE											
λ_1	λ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	3.527	2.634	2.634	2.630	2.630	2.630	2.630	2.630	2.630	2.635
0	0.2	3.544	2.657	2.657	2.652	2.652	2.651	2.651	2.653	2.652	2.658
0	0.5	3.671	2.829	2.833	2.824	2.824	2.825	2.825	2.827	2.826	2.834
0	0.8	4.469	3.829	3.870	3.821	3.823	3.835	3.834	3.826	3.838	3.837
0.2	0	3.529	2.634	2.634	2.630	2.630	2.630	2.630	2.630	2.630	2.635
0.2	0.2	3.547	2.657	2.657	2.652	2.652	2.652	2.652	2.653	2.652	2.658
0.2	0.5	3.674	2.831	2.833	2.827	2.826	2.825	2.825	2.828	2.826	2.834
0.2	0.8	4.473	3.837	3.871	3.829	3.826	3.835	3.834	3.832	3.839	3.844
0.5	0	3.599	2.636	2.635	2.631	2.631	2.631	2.631	2.629	2.631	2.634
0.5	0.2	3.616	2.659	2.658	2.655	2.654	2.653	2.653	2.652	2.653	2.657
0.5	0.5	3.743	2.839	2.835	2.834	2.832	2.826	2.826	2.829	2.827	2.834
0.5	0.8	4.532	3.867	3.873	3.859	3.838	3.836	3.835	3.844	3.842	3.854
0.8	0	4.002	2.639	2.639	2.634	2.634	2.634	2.634	2.629	2.634	2.634
0.8	0.2	4.018	2.665	2.661	2.660	2.659	2.656	2.656	2.652	2.656	2.657
0.8	0.5	4.134	2.860	2.839	2.854	2.845	2.831	2.831	2.828	2.831	2.834
0.8	0.8	4.870	4.006	3.879	4.000	3.865	3.842	3.842	3.848	3.848	3.856

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 2: RMSE and MAE of Spatial Panel Data Predictors: SAR and $\rho = 0.2$

RMSE											
λ_1	λ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	4.445	3.279	3.279	3.342	3.342	3.342	3.342	3.343	3.342	3.280
0	0.2	4.467	3.305	3.306	3.369	3.369	3.369	3.369	3.370	3.369	3.308
0	0.5	4.629	3.512	3.517	3.581	3.580	3.582	3.582	3.583	3.582	3.518
0	0.8	5.605	4.666	4.724	4.760	4.764	4.783	4.781	4.767	4.783	4.673
0.2	0	4.447	3.279	3.279	3.342	3.342	3.342	3.342	3.342	3.342	3.280
0.2	0.2	4.469	3.306	3.306	3.370	3.370	3.369	3.369	3.370	3.369	3.308
0.2	0.5	4.632	3.515	3.517	3.583	3.582	3.582	3.582	3.585	3.582	3.519
0.2	0.8	5.607	4.674	4.724	4.769	4.768	4.783	4.781	4.774	4.783	4.680
0.5	0	4.530	3.281	3.280	3.343	3.343	3.343	3.343	3.341	3.343	3.279
0.5	0.2	4.551	3.310	3.307	3.373	3.372	3.370	3.370	3.370	3.371	3.307
0.5	0.5	4.710	3.526	3.519	3.592	3.589	3.584	3.584	3.586	3.584	3.519
0.5	0.8	5.672	4.703	4.727	4.800	4.779	4.784	4.783	4.787	4.786	4.695
0.8	0	5.001	3.286	3.285	3.349	3.349	3.348	3.348	3.340	3.348	3.278
0.8	0.2	5.020	3.318	3.313	3.380	3.379	3.375	3.375	3.369	3.376	3.306
0.8	0.5	5.164	3.564	3.525	3.614	3.605	3.589	3.589	3.586	3.590	3.519
0.8	0.8	6.055	4.844	4.734	4.950	4.809	4.790	4.789	4.796	4.794	4.705
MAE											
λ_1	λ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	3.564	2.626	2.625	2.674	2.674	2.674	2.674	2.675	2.674	2.626
0	0.2	3.582	2.648	2.648	2.697	2.697	2.697	2.697	2.698	2.697	2.649
0	0.5	3.714	2.819	2.823	2.871	2.871	2.873	2.873	2.874	2.873	2.824
0	0.8	4.541	3.813	3.858	3.890	3.892	3.907	3.906	3.896	3.906	3.820
0.2	0	3.567	2.626	2.626	2.674	2.674	2.674	2.674	2.674	2.674	2.626
0.2	0.2	3.585	2.648	2.648	2.697	2.697	2.697	2.697	2.698	2.697	2.649
0.2	0.5	3.717	2.822	2.824	2.873	2.873	2.873	2.873	2.875	2.873	2.825
0.2	0.8	4.544	3.821	3.859	3.898	3.896	3.907	3.906	3.902	3.907	3.826
0.5	0	3.636	2.627	2.627	2.675	2.675	2.675	2.675	2.674	2.675	2.626
0.5	0.2	3.654	2.651	2.650	2.700	2.700	2.698	2.698	2.698	2.698	2.649
0.5	0.5	3.785	2.831	2.826	2.881	2.879	2.874	2.874	2.876	2.875	2.825
0.5	0.8	4.603	3.847	3.862	3.925	3.906	3.908	3.907	3.912	3.910	3.838
0.8	0	4.035	2.631	2.631	2.680	2.680	2.679	2.679	2.673	2.680	2.625
0.8	0.2	4.053	2.658	2.654	2.706	2.705	2.702	2.702	2.697	2.703	2.648
0.8	0.5	4.174	2.863	2.831	2.900	2.892	2.879	2.878	2.876	2.880	2.825
0.8	0.8	4.934	3.970	3.869	4.055	3.932	3.913	3.913	3.919	3.917	3.846

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 3: RMSE and MAE of Spatial Panel Data Predictors: SAR and $\rho = 0.5$

RMSE											
λ_1	λ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	4.760	3.250	3.250	3.715	3.715	3.715	3.715	3.716	3.715	3.251
0	0.2	4.785	3.276	3.276	3.745	3.745	3.745	3.745	3.745	3.745	3.279
0	0.5	4.973	3.480	3.487	3.975	3.974	3.980	3.979	3.977	3.977	3.484
0	0.8	6.091	4.625	4.687	5.264	5.271	5.302	5.298	5.278	5.289	4.629
0.2	0	4.763	3.250	3.250	3.716	3.716	3.716	3.716	3.716	3.716	3.251
0.2	0.2	4.787	3.277	3.277	3.745	3.745	3.745	3.745	3.746	3.745	3.279
0.2	0.5	4.976	3.483	3.487	3.977	3.976	3.979	3.979	3.979	3.977	3.486
0.2	0.8	6.093	4.631	4.687	5.272	5.274	5.302	5.298	5.284	5.290	4.634
0.5	0	4.841	3.252	3.252	3.718	3.718	3.718	3.718	3.715	3.718	3.250
0.5	0.2	4.865	3.282	3.279	3.749	3.748	3.747	3.746	3.746	3.747	3.278
0.5	0.5	5.050	3.495	3.489	3.984	3.982	3.980	3.980	3.981	3.980	3.489
0.5	0.8	6.155	4.649	4.691	5.299	5.284	5.302	5.299	5.297	5.294	4.646
0.8	0	5.287	3.261	3.259	3.726	3.726	3.724	3.724	3.715	3.725	3.249
0.8	0.2	5.310	3.299	3.287	3.758	3.758	3.753	3.753	3.746	3.755	3.277
0.8	0.5	5.479	3.548	3.498	4.008	4.000	3.986	3.986	3.984	3.988	3.490
0.8	0.8	6.514	4.730	4.703	5.426	5.315	5.307	5.306	5.310	5.306	4.666
MAE											
λ_1	λ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	3.818	2.603	2.603	2.975	2.975	2.975	2.975	2.975	2.975	2.604
0	0.2	3.839	2.625	2.625	2.998	2.998	2.999	2.999	2.999	2.998	2.627
0	0.5	3.994	2.793	2.799	3.187	3.187	3.191	3.191	3.189	3.190	2.796
0	0.8	4.941	3.780	3.828	4.297	4.302	4.327	4.324	4.307	4.315	3.784
0.2	0	3.821	2.603	2.603	2.975	2.975	2.975	2.975	2.975	2.975	2.604
0.2	0.2	3.842	2.626	2.626	2.999	2.999	2.999	2.999	2.999	2.999	2.627
0.2	0.5	3.998	2.796	2.799	3.189	3.188	3.191	3.191	3.190	3.190	2.798
0.2	0.8	4.945	3.785	3.829	4.303	4.304	4.326	4.323	4.313	4.316	3.787
0.5	0	3.887	2.605	2.605	2.977	2.977	2.976	2.976	2.975	2.977	2.603
0.5	0.2	3.908	2.630	2.627	3.001	3.001	3.000	3.000	2.999	3.000	2.626
0.5	0.5	4.062	2.806	2.802	3.195	3.194	3.192	3.192	3.193	3.192	2.801
0.5	0.8	5.000	3.802	3.833	4.326	4.313	4.327	4.324	4.323	4.319	3.798
0.8	0	4.263	2.612	2.611	2.983	2.983	2.982	2.982	2.974	2.983	2.602
0.8	0.2	4.283	2.644	2.634	3.009	3.009	3.005	3.005	3.000	3.006	2.626
0.8	0.5	4.428	2.852	2.809	3.215	3.208	3.197	3.197	3.195	3.199	2.801
0.8	0.8	5.309	3.872	3.844	4.435	4.339	4.331	4.330	4.335	4.330	3.815

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 4: RMSE and MAE of Spatial Panel Data Predictors: SAR and $\rho = 0.8$

RMSE											
λ_1	λ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	6.056	3.190	3.190	4.729	4.729	4.729	4.729	4.729	4.729	3.191
0	0.2	6.086	3.216	3.216	4.763	4.763	4.764	4.764	4.763	4.763	3.217
0	0.5	6.351	3.418	3.423	5.043	5.043	5.053	5.053	5.049	5.049	3.419
0	0.8	7.898	4.554	4.587	6.649	6.655	6.694	6.689	6.676	6.673	4.554
0.2	0	6.059	3.190	3.190	4.729	4.729	4.729	4.729	4.729	4.729	3.191
0.2	0.2	6.088	3.216	3.216	4.763	4.763	4.764	4.764	4.763	4.763	3.218
0.2	0.5	6.354	3.419	3.423	5.044	5.044	5.052	5.052	5.050	5.049	3.420
0.2	0.8	7.900	4.556	4.588	6.653	6.656	6.693	6.688	6.677	6.673	4.556
0.5	0	6.121	3.192	3.192	4.731	4.731	4.731	4.731	4.730	4.731	3.191
0.5	0.2	6.151	3.219	3.219	4.765	4.765	4.765	4.765	4.765	4.765	3.218
0.5	0.5	6.414	3.423	3.425	5.048	5.048	5.053	5.053	5.052	5.051	3.423
0.5	0.8	7.951	4.560	4.591	6.667	6.661	6.693	6.688	6.682	6.676	4.560
0.8	0	6.487	3.203	3.200	4.738	4.738	4.738	4.738	4.731	4.738	3.191
0.8	0.2	6.515	3.234	3.227	4.774	4.774	4.770	4.770	4.766	4.772	3.221
0.8	0.5	6.764	3.441	3.435	5.065	5.061	5.057	5.057	5.057	5.059	3.432
0.8	0.8	8.242	4.579	4.604	6.726	6.680	6.695	6.692	6.695	6.686	4.573
MAE											
λ_1	λ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	4.864	2.555	2.556	3.787	3.787	3.788	3.788	3.787	3.788	2.556
0	0.2	4.888	2.577	2.578	3.814	3.814	3.815	3.815	3.814	3.815	2.578
0	0.5	5.105	2.744	2.748	4.042	4.042	4.050	4.050	4.046	4.047	2.745
0	0.8	6.404	3.723	3.747	5.419	5.422	5.452	5.449	5.438	5.437	3.722
0.2	0	4.867	2.556	2.556	3.787	3.787	3.788	3.788	3.787	3.788	2.556
0.2	0.2	4.891	2.578	2.578	3.814	3.814	3.815	3.815	3.815	3.815	2.578
0.2	0.5	5.108	2.745	2.748	4.043	4.043	4.050	4.050	4.047	4.047	2.745
0.2	0.8	6.408	3.724	3.748	5.422	5.423	5.452	5.448	5.439	5.437	3.723
0.5	0	4.918	2.558	2.558	3.789	3.789	3.789	3.789	3.788	3.789	2.557
0.5	0.2	4.944	2.580	2.580	3.816	3.816	3.816	3.816	3.816	3.816	2.579
0.5	0.5	5.159	2.749	2.750	4.046	4.046	4.050	4.050	4.050	4.049	2.749
0.5	0.8	6.453	3.728	3.750	5.434	5.427	5.452	5.448	5.443	5.439	3.727
0.8	0	5.219	2.566	2.565	3.794	3.794	3.794	3.794	3.789	3.795	2.556
0.8	0.2	5.244	2.592	2.587	3.822	3.822	3.820	3.820	3.817	3.822	2.581
0.8	0.5	5.453	2.764	2.758	4.060	4.057	4.053	4.053	4.054	4.055	2.756
0.8	0.8	6.706	3.745	3.761	5.485	5.444	5.454	5.452	5.455	5.448	3.738

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 5: RMSE and MAE of Spatial Panel Data Predictors: SMA and $\rho = 0$

RMSE											
δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	4.396	3.290	3.290	3.284	3.284	3.284	3.284	3.285	3.284	3.291
0	0.2	4.405	3.299	3.299	3.293	3.293	3.293	3.293	3.295	3.293	3.301
0	0.5	4.435	3.336	3.337	3.330	3.330	3.330	3.330	3.332	3.330	3.339
0	0.8	4.484	3.399	3.403	3.393	3.392	3.393	3.393	3.396	3.394	3.404
0.2	0	4.390	3.290	3.290	3.284	3.284	3.284	3.284	3.285	3.284	3.291
0.2	0.2	4.399	3.299	3.299	3.294	3.294	3.293	3.293	3.295	3.293	3.301
0.2	0.5	4.428	3.337	3.337	3.331	3.331	3.330	3.330	3.332	3.330	3.339
0.2	0.8	4.478	3.401	3.403	3.395	3.394	3.393	3.393	3.396	3.394	3.404
0.5	0	4.393	3.290	3.290	3.285	3.285	3.285	3.285	3.284	3.285	3.290
0.5	0.2	4.402	3.300	3.299	3.295	3.295	3.293	3.293	3.294	3.294	3.300
0.5	0.5	4.431	3.339	3.338	3.334	3.333	3.330	3.330	3.332	3.331	3.338
0.5	0.8	4.481	3.405	3.403	3.399	3.399	3.394	3.393	3.397	3.395	3.404
0.8	0	4.412	3.291	3.291	3.285	3.285	3.285	3.285	3.283	3.285	3.289
0.8	0.2	4.420	3.302	3.300	3.296	3.296	3.294	3.294	3.293	3.294	3.299
0.8	0.5	4.449	3.342	3.338	3.336	3.336	3.331	3.331	3.331	3.331	3.337
0.8	0.8	4.498	3.411	3.404	3.405	3.403	3.394	3.394	3.397	3.395	3.403
MAE											
δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	3.527	2.634	2.634	2.630	2.630	2.630	2.630	2.630	2.630	2.635
0	0.2	3.534	2.642	2.643	2.637	2.637	2.637	2.637	2.639	2.637	2.643
0	0.5	3.558	2.674	2.676	2.670	2.669	2.669	2.669	2.671	2.670	2.677
0	0.8	3.599	2.728	2.732	2.723	2.723	2.724	2.724	2.726	2.725	2.732
0.2	0	3.522	2.634	2.634	2.630	2.630	2.630	2.630	2.630	2.630	2.635
0.2	0.2	3.530	2.643	2.643	2.638	2.638	2.637	2.637	2.638	2.638	2.643
0.2	0.5	3.554	2.675	2.676	2.671	2.671	2.669	2.669	2.672	2.670	2.677
0.2	0.8	3.595	2.730	2.732	2.725	2.725	2.724	2.724	2.727	2.725	2.732
0.5	0	3.526	2.635	2.635	2.630	2.630	2.630	2.630	2.629	2.630	2.634
0.5	0.2	3.533	2.644	2.643	2.639	2.639	2.638	2.638	2.638	2.638	2.643
0.5	0.5	3.558	2.677	2.676	2.673	2.673	2.670	2.670	2.671	2.670	2.676
0.5	0.8	3.600	2.734	2.732	2.729	2.728	2.724	2.724	2.727	2.725	2.732
0.8	0	3.542	2.635	2.635	2.630	2.630	2.630	2.630	2.629	2.630	2.633
0.8	0.2	3.550	2.645	2.643	2.640	2.640	2.638	2.638	2.637	2.638	2.642
0.8	0.5	3.574	2.680	2.677	2.675	2.675	2.670	2.670	2.671	2.671	2.675
0.8	0.8	3.616	2.738	2.733	2.733	2.732	2.725	2.725	2.727	2.726	2.732

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 6: RMSE and MAE of Spatial Panel Data Predictors: SMA and $\rho = 0.2$

RMSE											
δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	4.445	3.279	3.279	3.342	3.342	3.342	3.342	3.343	3.342	3.280
0	0.2	4.454	3.288	3.288	3.351	3.351	3.351	3.351	3.352	3.351	3.290
0	0.5	4.485	3.324	3.326	3.389	3.389	3.389	3.389	3.390	3.389	3.328
0	0.8	4.537	3.387	3.392	3.453	3.453	3.455	3.455	3.455	3.455	3.392
0.2	0	4.439	3.279	3.279	3.342	3.342	3.342	3.342	3.342	3.342	3.280
0.2	0.2	4.448	3.288	3.288	3.352	3.352	3.351	3.351	3.352	3.351	3.290
0.2	0.5	4.479	3.326	3.326	3.390	3.390	3.389	3.389	3.391	3.389	3.328
0.2	0.8	4.531	3.389	3.392	3.455	3.455	3.455	3.455	3.456	3.455	3.393
0.5	0	4.442	3.279	3.279	3.342	3.342	3.342	3.342	3.341	3.342	3.279
0.5	0.2	4.451	3.290	3.288	3.353	3.352	3.351	3.351	3.352	3.351	3.289
0.5	0.5	4.482	3.329	3.327	3.392	3.392	3.389	3.389	3.391	3.390	3.327
0.5	0.8	4.534	3.394	3.392	3.459	3.458	3.455	3.455	3.457	3.455	3.393
0.8	0	4.460	3.280	3.279	3.343	3.343	3.342	3.342	3.340	3.342	3.277
0.8	0.2	4.469	3.291	3.289	3.354	3.354	3.352	3.352	3.351	3.352	3.287
0.8	0.5	4.500	3.332	3.327	3.395	3.394	3.390	3.390	3.390	3.390	3.327
0.8	0.8	4.551	3.401	3.393	3.464	3.462	3.455	3.455	3.457	3.456	3.392
MAE											
δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	3.564	2.626	2.625	2.674	2.674	2.674	2.674	2.675	2.674	2.626
0	0.2	3.571	2.634	2.634	2.682	2.682	2.682	2.682	2.683	2.682	2.635
0	0.5	3.597	2.665	2.667	2.715	2.715	2.715	2.715	2.716	2.715	2.668
0	0.8	3.640	2.718	2.723	2.769	2.769	2.771	2.771	2.771	2.771	2.723
0.2	0	3.560	2.626	2.626	2.674	2.674	2.674	2.674	2.674	2.674	2.626
0.2	0.2	3.567	2.634	2.634	2.683	2.683	2.682	2.682	2.683	2.682	2.635
0.2	0.5	3.593	2.666	2.667	2.716	2.715	2.715	2.715	2.716	2.715	2.668
0.2	0.8	3.636	2.721	2.723	2.771	2.771	2.771	2.771	2.772	2.771	2.723
0.5	0	3.564	2.626	2.626	2.674	2.674	2.674	2.674	2.674	2.674	2.625
0.5	0.2	3.572	2.635	2.634	2.683	2.683	2.682	2.682	2.683	2.683	2.634
0.5	0.5	3.598	2.669	2.667	2.717	2.717	2.715	2.715	2.716	2.716	2.667
0.5	0.8	3.641	2.725	2.724	2.774	2.774	2.771	2.771	2.773	2.771	2.723
0.8	0	3.580	2.626	2.626	2.674	2.674	2.674	2.674	2.673	2.674	2.624
0.8	0.2	3.588	2.636	2.635	2.684	2.684	2.683	2.683	2.682	2.683	2.633
0.8	0.5	3.614	2.672	2.668	2.720	2.719	2.715	2.715	2.716	2.716	2.667
0.8	0.8	3.657	2.730	2.724	2.778	2.777	2.771	2.771	2.773	2.772	2.723

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 7: RMSE and MAE of Spatial Panel Data Predictors: SMA and $\rho = 0.5$

RMSE											
δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	4.760	3.250	3.250	3.715	3.715	3.715	3.715	3.716	3.715	3.251
0	0.2	4.769	3.259	3.259	3.725	3.725	3.725	3.725	3.726	3.725	3.261
0	0.5	4.805	3.295	3.297	3.765	3.765	3.767	3.767	3.766	3.766	3.298
0	0.8	4.864	3.356	3.363	3.834	3.833	3.839	3.839	3.836	3.837	3.360
0.2	0	4.754	3.250	3.250	3.715	3.715	3.715	3.715	3.716	3.715	3.251
0.2	0.2	4.764	3.260	3.259	3.726	3.725	3.725	3.725	3.726	3.725	3.261
0.2	0.5	4.799	3.296	3.297	3.766	3.766	3.767	3.767	3.767	3.766	3.299
0.2	0.8	4.858	3.359	3.363	3.835	3.834	3.839	3.839	3.837	3.837	3.362
0.5	0	4.758	3.250	3.250	3.716	3.716	3.716	3.716	3.715	3.716	3.249
0.5	0.2	4.768	3.261	3.260	3.726	3.726	3.725	3.725	3.725	3.725	3.260
0.5	0.5	4.803	3.300	3.298	3.768	3.767	3.766	3.766	3.767	3.766	3.299
0.5	0.8	4.862	3.364	3.363	3.838	3.837	3.838	3.838	3.838	3.837	3.363
0.8	0	4.776	3.251	3.251	3.716	3.716	3.716	3.716	3.713	3.716	3.248
0.8	0.2	4.785	3.263	3.260	3.727	3.727	3.725	3.725	3.724	3.726	3.259
0.8	0.5	4.820	3.304	3.299	3.770	3.769	3.766	3.766	3.767	3.767	3.298
0.8	0.8	4.878	3.370	3.364	3.842	3.841	3.838	3.838	3.839	3.838	3.363
MAE											
δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	3.818	2.603	2.603	2.975	2.975	2.975	2.975	2.975	2.975	2.604
0	0.2	3.826	2.611	2.611	2.983	2.983	2.983	2.983	2.983	2.983	2.613
0	0.5	3.856	2.641	2.643	3.016	3.015	3.017	3.017	3.017	3.017	2.644
0	0.8	3.906	2.694	2.699	3.074	3.073	3.078	3.078	3.075	3.077	2.697
0.2	0	3.815	2.603	2.603	2.975	2.975	2.975	2.975	2.975	2.975	2.604
0.2	0.2	3.823	2.611	2.611	2.983	2.983	2.983	2.983	2.983	2.983	2.613
0.2	0.5	3.853	2.643	2.644	3.016	3.016	3.017	3.017	3.017	3.017	2.645
0.2	0.8	3.903	2.696	2.700	3.075	3.075	3.078	3.078	3.077	3.077	2.699
0.5	0	3.819	2.603	2.603	2.975	2.975	2.975	2.975	2.974	2.975	2.602
0.5	0.2	3.827	2.613	2.612	2.983	2.983	2.983	2.983	2.983	2.983	2.612
0.5	0.5	3.857	2.645	2.644	3.018	3.018	3.017	3.017	3.017	3.017	2.645
0.5	0.8	3.908	2.700	2.700	3.078	3.077	3.078	3.078	3.078	3.077	2.700
0.8	0	3.835	2.604	2.604	2.975	2.975	2.975	2.975	2.973	2.975	2.601
0.8	0.2	3.843	2.614	2.612	2.984	2.984	2.983	2.983	2.982	2.983	2.610
0.8	0.5	3.873	2.649	2.645	3.020	3.019	3.017	3.017	3.017	3.018	2.644
0.8	0.8	3.923	2.706	2.701	3.081	3.080	3.078	3.078	3.078	3.078	2.699

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 8: RMSE and MAE of Spatial Panel Data Predictors: SMA and $\rho = 0.8$

RMSE											
δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	6.056	3.190	3.190	4.729	4.729	4.729	4.729	4.729	4.729	3.191
0	0.2	6.063	3.199	3.199	4.738	4.738	4.739	4.739	4.739	4.739	3.200
0	0.5	6.104	3.235	3.237	4.784	4.783	4.788	4.788	4.786	4.786	3.236
0	0.8	6.182	3.297	3.301	4.865	4.864	4.875	4.875	4.870	4.871	3.297
0.2	0	6.052	3.190	3.190	4.729	4.729	4.729	4.729	4.729	4.729	3.191
0.2	0.2	6.059	3.199	3.199	4.738	4.738	4.739	4.739	4.739	4.739	3.200
0.2	0.5	6.100	3.236	3.237	4.784	4.783	4.787	4.787	4.786	4.786	3.236
0.2	0.8	6.178	3.298	3.301	4.865	4.864	4.874	4.874	4.871	4.871	3.298
0.5	0	6.055	3.190	3.190	4.729	4.729	4.729	4.729	4.729	4.729	3.190
0.5	0.2	6.062	3.200	3.200	4.739	4.738	4.739	4.739	4.739	4.739	3.200
0.5	0.5	6.103	3.237	3.237	4.785	4.784	4.787	4.787	4.786	4.786	3.237
0.5	0.8	6.181	3.299	3.302	4.867	4.866	4.874	4.874	4.872	4.871	3.299
0.8	0	6.069	3.191	3.191	4.729	4.729	4.730	4.730	4.728	4.730	3.190
0.8	0.2	6.076	3.201	3.200	4.739	4.739	4.739	4.739	4.739	4.739	3.200
0.8	0.5	6.117	3.238	3.238	4.786	4.785	4.787	4.787	4.787	4.786	3.238
0.8	0.8	6.195	3.301	3.302	4.868	4.867	4.873	4.873	4.872	4.871	3.300
MAE											
δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General
0	0	4.864	2.555	2.556	3.787	3.787	3.788	3.788	3.787	3.788	2.556
0	0.2	4.870	2.563	2.564	3.794	3.794	3.795	3.795	3.795	3.795	2.564
0	0.5	4.904	2.594	2.596	3.831	3.831	3.835	3.835	3.833	3.833	2.594
0	0.8	4.969	2.647	2.650	3.898	3.898	3.906	3.906	3.902	3.904	2.647
0.2	0	4.861	2.556	2.556	3.787	3.787	3.788	3.788	3.787	3.787	2.556
0.2	0.2	4.867	2.564	2.564	3.794	3.794	3.795	3.795	3.795	3.795	2.565
0.2	0.5	4.901	2.594	2.596	3.831	3.831	3.834	3.834	3.833	3.833	2.595
0.2	0.8	4.966	2.647	2.650	3.898	3.898	3.906	3.906	3.903	3.904	2.647
0.5	0	4.865	2.556	2.556	3.787	3.787	3.788	3.788	3.787	3.788	2.556
0.5	0.2	4.871	2.564	2.564	3.795	3.794	3.795	3.795	3.795	3.795	2.564
0.5	0.5	4.905	2.596	2.596	3.832	3.831	3.834	3.834	3.834	3.833	2.595
0.5	0.8	4.970	2.649	2.651	3.899	3.899	3.905	3.905	3.904	3.904	2.648
0.8	0	4.877	2.556	2.557	3.787	3.787	3.788	3.788	3.787	3.788	2.556
0.8	0.2	4.883	2.566	2.565	3.795	3.795	3.795	3.795	3.795	3.795	2.564
0.8	0.5	4.918	2.597	2.597	3.833	3.832	3.834	3.834	3.834	3.834	2.596
0.8	0.8	4.982	2.650	2.652	3.901	3.900	3.905	3.905	3.904	3.904	2.650

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 9: RMSE and MAE of Spatial Panel Data Predictors: SARMA and $\rho = 0$

RMSE														
λ_1	λ_2	δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General	
0.2	0.2	0.2	0.2	4.464	3.363	3.361	3.357	3.356	3.354	3.354	3.356	3.354	3.362	
0.2	0.2	0.2	0.5	4.535	3.458	3.457	3.452	3.451	3.447	3.447	3.450	3.448	3.457	
0.2	0.2	0.5	0.2	4.495	3.366	3.362	3.360	3.359	3.355	3.354	3.355	3.355	3.361	
0.2	0.2	0.5	0.5	4.565	3.464	3.457	3.458	3.456	3.447	3.447	3.450	3.448	3.456	
0.2	0.5	0.2	0.2	4.695	3.674	3.674	3.667	3.664	3.660	3.660	3.665	3.662	3.673	
0.2	0.5	0.2	0.5	4.892	3.927	3.934	3.920	3.913	3.913	3.913	3.918	3.915	3.928	
0.2	0.5	0.5	0.2	4.724	3.683	3.675	3.676	3.671	3.661	3.661	3.665	3.663	3.673	
0.2	0.5	0.5	0.5	4.920	3.942	3.935	3.935	3.920	3.914	3.913	3.919	3.916	3.928	
0.5	0.2	0.2	0.2	4.592	3.368	3.363	3.363	3.362	3.356	3.356	3.355	3.356	3.361	
0.5	0.2	0.2	0.5	4.661	3.470	3.459	3.464	3.460	3.449	3.448	3.450	3.450	3.457	
0.5	0.2	0.5	0.2	4.703	3.372	3.364	3.367	3.365	3.357	3.357	3.355	3.357	3.361	
0.5	0.2	0.5	0.5	4.770	3.479	3.460	3.473	3.467	3.450	3.450	3.450	3.451	3.456	
0.5	0.5	0.2	0.2	4.817	3.693	3.676	3.686	3.677	3.663	3.662	3.666	3.664	3.673	
0.5	0.5	0.2	0.5	5.009	3.959	3.936	3.951	3.926	3.915	3.915	3.921	3.918	3.929	
0.5	0.5	0.5	0.2	4.922	3.709	3.678	3.702	3.685	3.664	3.664	3.665	3.666	3.673	
0.5	0.5	0.5	0.5	5.110	3.988	3.938	3.980	3.936	3.917	3.916	3.921	3.919	3.929	
MAE														
λ_1	λ_2	δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General	
0.2	0.2	0.2	0.2	3.584	2.695	2.695	2.691	2.691	2.688	2.688	2.690	2.689	2.695	
0.2	0.2	0.2	0.5	3.643	2.776	2.775	2.771	2.770	2.767	2.767	2.770	2.768	2.775	
0.2	0.2	0.5	0.2	3.611	2.698	2.695	2.693	2.693	2.689	2.689	2.689	2.689	2.694	
0.2	0.2	0.5	0.5	3.670	2.781	2.776	2.776	2.775	2.768	2.767	2.770	2.769	2.775	
0.2	0.5	0.2	0.2	3.775	2.955	2.956	2.950	2.947	2.945	2.944	2.948	2.946	2.955	
0.2	0.5	0.2	0.5	3.940	3.173	3.176	3.167	3.160	3.160	3.160	3.165	3.162	3.173	
0.2	0.5	0.5	0.2	3.802	2.963	2.956	2.958	2.953	2.945	2.945	2.948	2.947	2.954	
0.2	0.5	0.5	0.5	3.967	3.186	3.178	3.180	3.167	3.161	3.160	3.166	3.163	3.172	
0.5	0.2	0.2	0.2	3.693	2.700	2.696	2.696	2.695	2.690	2.690	2.689	2.690	2.694	
0.5	0.2	0.2	0.5	3.751	2.786	2.777	2.781	2.778	2.769	2.768	2.770	2.770	2.775	
0.5	0.2	0.5	0.2	3.788	2.704	2.697	2.699	2.698	2.691	2.691	2.689	2.691	2.694	
0.5	0.2	0.5	0.5	3.844	2.794	2.778	2.789	2.784	2.770	2.770	2.769	2.771	2.775	
0.5	0.5	0.2	0.2	3.880	2.972	2.958	2.966	2.958	2.946	2.946	2.949	2.948	2.955	
0.5	0.5	0.2	0.5	4.043	3.201	3.179	3.195	3.172	3.162	3.161	3.167	3.164	3.173	
0.5	0.5	0.5	0.2	3.971	2.986	2.959	2.980	2.966	2.948	2.948	2.949	2.949	2.954	
0.5	0.5	0.5	0.5	4.131	3.227	3.181	3.221	3.181	3.163	3.163	3.167	3.166	3.173	

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 10: RMSE and MAE of Spatial Panel Data Predictors: SARMA and $\rho = 0.2$

RMSE														
λ_1	λ_2	δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General	
0.2	0.2	0.2	0.2	4.515	3.352	3.350	3.416	3.415	3.413	3.413	3.415	3.414	3.351	
0.2	0.2	0.2	0.5	4.589	3.447	3.445	3.513	3.511	3.509	3.509	3.511	3.509	3.446	
0.2	0.2	0.5	0.2	4.546	3.355	3.351	3.419	3.418	3.414	3.414	3.414	3.414	3.350	
0.2	0.2	0.5	0.5	4.619	3.454	3.446	3.518	3.516	3.510	3.509	3.511	3.510	3.445	
0.2	0.5	0.2	0.2	4.754	3.661	3.662	3.731	3.729	3.727	3.727	3.730	3.727	3.661	
0.2	0.5	0.2	0.5	4.959	3.912	3.921	3.988	3.983	3.986	3.986	3.987	3.986	3.913	
0.2	0.5	0.5	0.2	4.783	3.671	3.663	3.740	3.735	3.728	3.728	3.731	3.728	3.661	
0.2	0.5	0.5	0.5	4.986	3.926	3.922	4.002	3.989	3.986	3.986	3.989	3.987	3.915	
0.5	0.2	0.2	0.2	4.642	3.359	3.352	3.421	3.421	3.415	3.415	3.415	3.416	3.351	
0.5	0.2	0.2	0.5	4.714	3.461	3.448	3.524	3.521	3.511	3.511	3.512	3.512	3.446	
0.5	0.2	0.5	0.2	4.752	3.365	3.354	3.425	3.424	3.417	3.417	3.414	3.417	3.350	
0.5	0.2	0.5	0.5	4.822	3.474	3.449	3.532	3.527	3.512	3.512	3.512	3.513	3.445	
0.5	0.5	0.2	0.2	4.875	3.683	3.665	3.749	3.740	3.729	3.729	3.732	3.730	3.662	
0.5	0.5	0.2	0.5	5.074	3.944	3.924	4.017	3.995	3.988	3.987	3.992	3.989	3.917	
0.5	0.5	0.5	0.2	4.979	3.703	3.667	3.763	3.748	3.731	3.731	3.732	3.732	3.661	
0.5	0.5	0.5	0.5	5.174	3.975	3.926	4.043	4.004	3.989	3.989	3.992	3.991	3.917	
MAE														
λ_1	λ_2	δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General	
0.2	0.2	0.2	0.2	3.624	2.687	2.686	2.736	2.736	2.734	2.734	2.735	2.734	2.686	
0.2	0.2	0.2	0.5	3.685	2.767	2.766	2.817	2.816	2.814	2.814	2.816	2.815	2.766	
0.2	0.2	0.5	0.2	3.651	2.690	2.686	2.738	2.738	2.734	2.734	2.735	2.735	2.685	
0.2	0.2	0.5	0.5	3.712	2.773	2.767	2.822	2.820	2.815	2.815	2.816	2.815	2.766	
0.2	0.5	0.2	0.2	3.821	2.945	2.946	2.999	2.996	2.995	2.995	2.997	2.996	2.945	
0.2	0.5	0.2	0.5	3.993	3.160	3.166	3.220	3.214	3.216	3.216	3.218	3.216	3.160	
0.2	0.5	0.5	0.2	3.848	2.954	2.947	3.006	3.002	2.995	2.995	2.998	2.996	2.945	
0.2	0.5	0.5	0.5	4.020	3.173	3.167	3.232	3.220	3.216	3.216	3.220	3.217	3.161	
0.5	0.2	0.2	0.2	3.731	2.693	2.688	2.741	2.740	2.736	2.735	2.735	2.736	2.686	
0.5	0.2	0.2	0.5	3.791	2.780	2.769	2.827	2.824	2.816	2.816	2.817	2.817	2.766	
0.5	0.2	0.5	0.2	3.826	2.698	2.689	2.744	2.743	2.737	2.737	2.735	2.738	2.685	
0.5	0.2	0.5	0.5	3.884	2.790	2.770	2.834	2.830	2.817	2.817	2.817	2.818	2.766	
0.5	0.5	0.2	0.2	3.925	2.964	2.948	3.014	3.006	2.997	2.997	2.999	2.998	2.945	
0.5	0.5	0.2	0.5	4.094	3.189	3.169	3.245	3.225	3.218	3.217	3.221	3.219	3.163	
0.5	0.5	0.5	0.2	4.015	2.982	2.950	3.026	3.014	2.998	2.998	2.999	2.999	2.945	
0.5	0.5	0.5	0.5	4.180	3.216	3.171	3.268	3.234	3.219	3.219	3.222	3.221	3.163	

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 11: RMSE and MAE of Spatial Panel Data Predictors: SARMA and $\rho = 0.5$

RMSE														
λ_1	λ_2	δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General	
0.2	0.2	0.2	0.2	4.837	3.323	3.321	3.794	3.794	3.793	3.793	3.794	3.793	3.323	
0.2	0.2	0.2	0.5	4.922	3.416	3.416	3.898	3.897	3.898	3.898	3.898	3.897	3.416	
0.2	0.2	0.5	0.2	4.867	3.327	3.322	3.797	3.796	3.793	3.793	3.794	3.794	3.321	
0.2	0.2	0.5	0.5	4.951	3.423	3.417	3.902	3.901	3.898	3.898	3.899	3.898	3.416	
0.2	0.5	0.2	0.2	5.115	3.626	3.631	4.136	4.134	4.139	4.138	4.137	4.136	3.627	
0.2	0.5	0.2	0.5	5.351	3.873	3.890	4.414	4.411	4.424	4.423	4.419	4.418	3.874	
0.2	0.5	0.5	0.2	5.143	3.635	3.633	4.143	4.139	4.139	4.138	4.139	4.137	3.629	
0.2	0.5	0.5	0.5	5.377	3.884	3.891	4.425	4.417	4.423	4.422	4.422	4.419	3.879	
0.5	0.2	0.2	0.2	4.958	3.333	3.324	3.800	3.799	3.795	3.795	3.795	3.796	3.322	
0.5	0.2	0.2	0.5	5.041	3.432	3.419	3.908	3.905	3.900	3.900	3.900	3.900	3.417	
0.5	0.2	0.5	0.2	5.062	3.343	3.326	3.805	3.803	3.797	3.796	3.794	3.798	3.321	
0.5	0.2	0.5	0.5	5.143	3.448	3.422	3.916	3.912	3.901	3.901	3.901	3.902	3.416	
0.5	0.5	0.2	0.2	5.229	3.646	3.635	4.151	4.144	4.140	4.140	4.141	4.139	3.631	
0.5	0.5	0.2	0.5	5.460	3.896	3.894	4.439	4.423	4.425	4.424	4.425	4.422	3.883	
0.5	0.5	0.5	0.2	5.328	3.666	3.637	4.163	4.152	4.141	4.141	4.142	4.142	3.631	
0.5	0.5	0.5	0.5	5.554	3.919	3.896	4.461	4.432	4.426	4.425	4.427	4.425	3.884	
MAE														
λ_1	λ_2	δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General	
0.2	0.2	0.2	0.2	3.885	2.663	2.662	3.039	3.039	3.038	3.038	3.038	3.038	2.663	
0.2	0.2	0.2	0.5	3.957	2.742	2.743	3.126	3.125	3.126	3.126	3.126	3.125	2.743	
0.2	0.2	0.5	0.2	3.911	2.668	2.663	3.041	3.040	3.038	3.038	3.038	3.039	2.662	
0.2	0.2	0.5	0.5	3.983	2.749	2.744	3.129	3.128	3.126	3.126	3.126	3.126	2.742	
0.2	0.5	0.2	0.2	4.116	2.917	2.921	3.322	3.321	3.325	3.324	3.323	3.322	2.918	
0.2	0.5	0.2	0.5	4.314	3.128	3.141	3.561	3.558	3.567	3.567	3.564	3.563	3.128	
0.2	0.5	0.5	0.2	4.141	2.926	2.923	3.329	3.325	3.325	3.324	3.325	3.323	2.919	
0.2	0.5	0.5	0.5	4.339	3.138	3.142	3.570	3.562	3.567	3.566	3.566	3.564	3.132	
0.5	0.2	0.2	0.2	3.987	2.672	2.665	3.044	3.043	3.040	3.040	3.039	3.041	2.663	
0.5	0.2	0.2	0.5	4.058	2.756	2.745	3.134	3.132	3.127	3.127	3.128	3.128	2.743	
0.5	0.2	0.5	0.2	4.075	2.681	2.667	3.047	3.046	3.041	3.041	3.039	3.042	2.662	
0.5	0.2	0.5	0.5	4.145	2.770	2.747	3.141	3.137	3.128	3.128	3.128	3.130	2.743	
0.5	0.5	0.2	0.2	4.214	2.935	2.925	3.336	3.330	3.326	3.326	3.327	3.325	2.921	
0.5	0.5	0.2	0.5	4.409	3.150	3.144	3.582	3.568	3.568	3.568	3.569	3.566	3.136	
0.5	0.5	0.5	0.2	4.298	2.952	2.927	3.346	3.337	3.327	3.327	3.328	3.328	2.921	
0.5	0.5	0.5	0.5	4.491	3.169	3.147	3.601	3.575	3.569	3.569	3.571	3.568	3.137	

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.

Table 12: RMSE and MAE of Spatial Panel Data Predictors: SARMA and $\rho = 0.8$

RMSE														
λ_1	λ_2	δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General	
0.2	0.2	0.2	0.2	6.148	3.260	3.260	4.819	4.819	4.821	4.821	4.821	4.820	3.260	
0.2	0.2	0.2	0.5	6.264	3.350	3.353	4.942	4.941	4.949	4.949	4.947	4.946	3.351	
0.2	0.2	0.5	0.2	6.172	3.262	3.261	4.821	4.820	4.821	4.821	4.821	4.821	3.261	
0.2	0.2	0.5	0.5	6.288	3.353	3.354	4.944	4.943	4.949	4.949	4.948	4.947	3.352	
0.2	0.5	0.2	0.2	6.545	3.557	3.564	5.235	5.235	5.249	5.248	5.245	5.244	3.557	
0.2	0.5	0.2	0.5	6.877	3.802	3.815	5.576	5.579	5.602	5.600	5.594	5.593	3.802	
0.2	0.5	0.5	0.2	6.567	3.560	3.565	5.238	5.237	5.248	5.248	5.246	5.244	3.559	
0.2	0.5	0.5	0.5	6.899	3.805	3.816	5.581	5.582	5.601	5.600	5.596	5.594	3.803	
0.5	0.2	0.2	0.2	6.246	3.265	3.263	4.823	4.823	4.823	4.823	4.823	4.823	3.262	
0.5	0.2	0.2	0.5	6.360	3.356	3.357	4.948	4.946	4.950	4.950	4.950	4.949	3.354	
0.5	0.2	0.5	0.2	6.330	3.269	3.265	4.826	4.826	4.824	4.823	4.823	4.825	3.263	
0.5	0.2	0.5	0.5	6.443	3.361	3.359	4.952	4.950	4.951	4.950	4.951	4.951	3.357	
0.5	0.5	0.2	0.2	6.637	3.564	3.567	5.243	5.241	5.249	5.249	5.249	5.247	3.562	
0.5	0.5	0.2	0.5	6.965	3.808	3.819	5.588	5.586	5.602	5.601	5.599	5.596	3.806	
0.5	0.5	0.5	0.2	6.716	3.569	3.570	5.249	5.246	5.250	5.249	5.250	5.249	3.565	
0.5	0.5	0.5	0.5	7.041	3.814	3.821	5.598	5.591	5.602	5.601	5.601	5.598	3.810	
MAE														
λ_1	λ_2	δ_1	δ_2	OLS	Case 1	Case 2	Case 3	Case 4	Case 5	Case 6	Case 7	Case 8	General	
0.2	0.2	0.2	0.2	4.941	2.614	2.614	3.860	3.859	3.862	3.861	3.861	3.861	2.614	
0.2	0.2	0.2	0.5	5.037	2.690	2.692	3.960	3.960	3.966	3.966	3.964	3.964	2.690	
0.2	0.2	0.5	0.2	4.962	2.615	2.615	3.861	3.860	3.861	3.861	3.862	3.861	2.614	
0.2	0.2	0.5	0.5	5.058	2.692	2.693	3.962	3.961	3.966	3.966	3.965	3.965	2.691	
0.2	0.5	0.2	0.2	5.267	2.863	2.868	4.202	4.203	4.213	4.213	4.210	4.209	2.862	
0.2	0.5	0.2	0.5	5.544	3.071	3.081	4.490	4.493	4.510	4.509	4.504	4.504	3.071	
0.2	0.5	0.5	0.2	5.287	2.865	2.869	4.205	4.205	4.213	4.212	4.211	4.210	2.864	
0.2	0.5	0.5	0.5	5.564	3.074	3.082	4.495	4.495	4.510	4.509	4.506	4.504	3.072	
0.5	0.2	0.2	0.2	5.023	2.618	2.617	3.863	3.862	3.862	3.862	3.863	3.863	2.615	
0.5	0.2	0.2	0.5	5.118	2.695	2.695	3.965	3.964	3.967	3.967	3.967	3.966	2.693	
0.5	0.2	0.5	0.2	5.094	2.622	2.619	3.865	3.864	3.863	3.863	3.863	3.864	2.616	
0.5	0.2	0.5	0.5	5.188	2.699	2.697	3.969	3.967	3.967	3.967	3.968	3.968	2.696	
0.5	0.5	0.2	0.2	5.345	2.868	2.871	4.209	4.208	4.214	4.213	4.213	4.212	2.866	
0.5	0.5	0.2	0.5	5.620	3.077	3.085	4.500	4.498	4.511	4.510	4.508	4.506	3.075	
0.5	0.5	0.5	0.2	5.414	2.872	2.873	4.214	4.211	4.214	4.214	4.215	4.213	2.869	
0.5	0.5	0.5	0.5	5.687	3.082	3.087	4.509	4.502	4.511	4.510	4.510	4.508	3.078	

Notes: $N = 49$, $T = 10$. 1,000 replications and 1 year forecasting ahead.